

Advanced Machine Learning Lecture 16

Beta Processes

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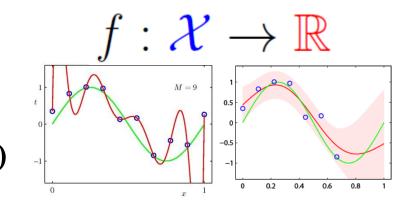


Announcement

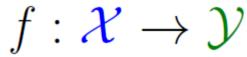
- Exercise sheet 3 will be made available tonight
 - Dirichlet Process Mixture Models
 - Gibbs Sampling
 - Finite Mixtures
 - DPMM Sampling

This Lecture: Advanced Machine Learning

- Regression Approaches
 - Linear Regression
 - Regularization (Ridge, Lasso)
 - Kernels (Kernel Ridge Regression)
 - Gaussian Processes



- Bayesian Estimation & Bayesian Non-Parametrics
 - Prob. Distributions, Approx. Inference
 - Mixture Models & EM
 - Dirichlet Processes
 - Latent Factor Models
 - Beta Processes
- SVMs and Structured Output Learning
 - SV Regression, SVDD
 - Large-margin Learning





Topics of This Lecture

- Latent Factor Models
 - Recap
- Towards Infinite Latent Factor Models
 - General formulation
 - Priors on binary matrices
 - Finite latent feature model
 - Left-ordered binary matrices
 - Indian Buffet Process
- Beta Processes
 - Properties
 - Stick-Breaking construction
 - Efficient Inference
 - Applications



Recap: Latent Factor Models

Mixture Models

- > Assume that each observation was generated by $\emph{exactly}$ one of K components.
- > The uncertainty is just about which component is responsible.

Latent Factor Models

- \triangleright Each observation is influenced by *each* of K components (factors or features) in a different way.
- Sparse factor models: only a small subset of factors is active for each observation.



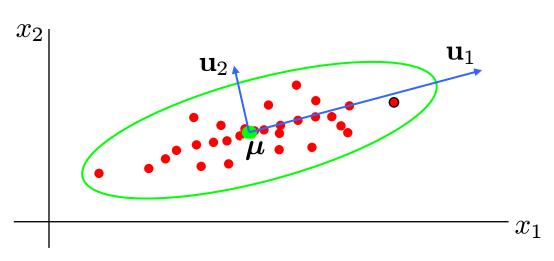
Recap: Principal Component Analysis

- Find the projection that maximizes the variance
 - Covariance matrix of the data

$$\mathbf{S} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \bar{\mathbf{x}}) (\mathbf{x}_n - \bar{\mathbf{x}})^T$$

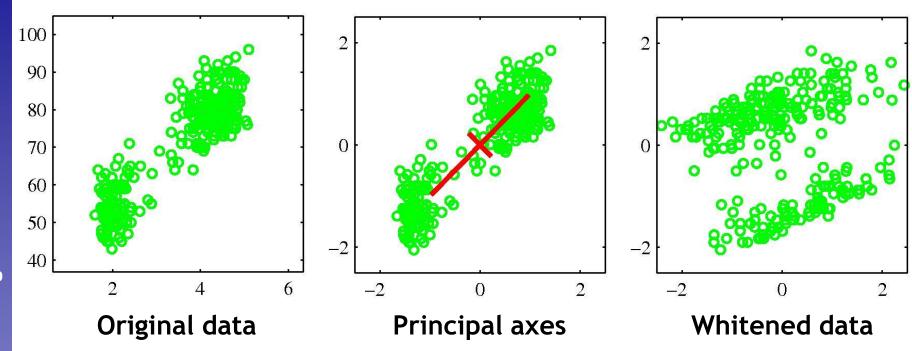
> Optimal linear projection into a K-dimensional space is given by the first K eigenvectors $\mathbf{u}_1,...,\mathbf{u}_K$ of \mathbf{S} .

$$\mathbf{y}_n = \mathbf{U}_{1..K} \mathbf{x}_n$$





Recap: PCA for Whitening



- Whitening procedure
 - Define for each data point the transformed value as

$$\mathbf{y}_n = \mathbf{L}^{-1/2} \mathbf{U}^T (\mathbf{x}_n - \bar{\mathbf{x}})$$
 $\mathbf{L} = \operatorname{diag}\{\lambda_i\}$ $\mathbf{U} = [\mathbf{u}_1, ..., \mathbf{u}_D]$

 \Rightarrow The transformed set $\{y_n\}$ has zero mean and unit covariance.



Recap: Probabilistic PCA

Graphical Model

- Introduce an explicit latent variable z corresponding to the principal component subspace.
- > Define a Gaussian prior distribution $p(\mathbf{z}) = \mathcal{N}(\mathbf{z}|\mathbf{0}, \mathbf{I})$
- Conditional distribution also Gaussian $p(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \sigma^2\mathbf{I})$
- Because of this linear-Gaussian model, the marginal distribution will also be Gaussian

$$p(\mathbf{x}) = \int p(\mathbf{x}|\mathbf{z})p(\mathbf{z})d\mathbf{z} = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \mathbf{C}), \qquad \mathbf{C} = \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I}$$

Posterior distribution (again Gaussian)

$$p(\mathbf{z}|\mathbf{x}) = \mathcal{N}\left(\mathbf{z}|\mathbf{M}^{-1}\mathbf{W}^{T}(\mathbf{x}-\boldsymbol{\mu}), \sigma^{2}\mathbf{M}^{-1}\right), \ \mathbf{M} = \mathbf{W}^{T}\mathbf{W} + \sigma^{2}\mathbf{I}$$

 \mathbf{X}_n



Recap: Interpretation of Probabilistic PCA

Analysis

- Marginal distribution: $p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|oldsymbol{\mu}, \mathbf{C}),$
- > Covariance matrix: $\mathbf{C} = \mathbf{W}\mathbf{W}^T + \sigma^2\mathbf{I}$
- \Rightarrow The columns of W define the principal subspace of PCA.
- Maximum Likelihood estimates

$$egin{aligned} oldsymbol{\mu}_{ ext{ML}} &= ar{\mathbf{x}} \ oldsymbol{W}_{ ext{ML}} &= oldsymbol{\mathbf{U}}_K (\mathbf{L}_K - \sigma^2 \mathbf{I})^{1/2} \mathbf{R} \ & \sigma_{ ext{ML}}^2 &= rac{1}{D-K} \sum_{i=K+1}^D \lambda_i \end{aligned}$$

 \Rightarrow The model correctly captures the variance of the data along the principal axes and approximates the variance in all remaining directions by σ^2 , the average of the discarded eigenvalues.



Recap: Examples of Latent Factor Models

- Probabilistic PCA (pPCA)
 - Linear-Gaussian model with isotropic covariance

$$p(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \sigma^2\mathbf{I})$$

- Factor Analysis (FA)
 - Same linear-Gaussian model, but with diagonal covariance

$$p(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \boldsymbol{\Psi}) \qquad \quad \boldsymbol{\Psi} = \text{diag}\{\psi_i\}$$

- Independent Component Analysis (ICA)
 - Observed variables are related linearly to the latent variables, but the latent distribution is non-Gaussian.
 - ightharpoonup Assumption: latent variables z_i are independent.

$$p(\mathbf{z}) = \prod_{j=1}^{K} p(z_j)$$



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Towards Infinite Latent Factor Models

- General formulation
- Priors on binary matrices
- Finite latent feature model
- Left-ordered binary matrices
- Indian Buffet Process

Beta Processes

- Properties
- Stick-Breaking construction
- Efficient Inference
- Applications

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Recap: General Latent Factor Models

General formulation

Assume that the data are generated by noisy weighted combination of latent factors

$$\mathbf{x}_n = \mathbf{F}\mathbf{y}_n + \boldsymbol{\epsilon}$$

- Mixture Models: DPs enforce that the main part of the probability mass is concentrated on few cluster components.
- Latent Factor Models: enforce that each object is represented by a sparse subset of an unbounded number of features.

Incorporating sparsity

- Decompose ${\bf F}$ into the product of two components: ${\bf F}={\bf Z}\otimes {\bf W}$, where \otimes is the Hadamard product (element-wise product).
 - z_{mk} is a binary mask variable indicating whether factor k is "on".
 - $-\ w_{mk}$ is a continuous weight variable.
- \Rightarrow Enforce sparsity by restricting the non-zero entries in \mathbb{Z} .

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Priors on Latent Factor Models

Defining suitable priors

- We will focus on defining a prior on ${\bf Z}$, since the effective dimensionality of the latent feature model is determined by ${\bf Z}$.
- Assuming that ${\bf Z}$ is sparse, we can define a prior for infinite latent feature models by defining a distribution over infinite binary matrices.

Desiderata for such a distribution

- Objects should be exchangeable.
- Inference should be tractable.

Procedure

Start with a model that assumes a finite number of features and consider the limit as this number approaches infinity.



- Modeling assumptions
 - $\,\,lacksquare$ We have N objects and K features.
 - ightharpoonup Binary variables z_{nk} indicates that object n possesses feature k.
 - > Each object possesses feature k with probability π_k and features are generated independently.
 - \Rightarrow The probability of a matrix ${\bf Z}$ given $\pi=\{\pi_1,...,\pi_k\}$ is given by a Binomial distribution

$$p(\mathbf{Z}|\boldsymbol{\pi}) = \prod_{k=1}^{K} \prod_{n=1}^{N} p(z_{nk}|\pi_k) = \prod_{k=1}^{K} \pi_k^{m_k} (1 - \pi_k)^{N - m_k}$$

where $m_k = \sum_{n=1}^N z_{nk}$ is the number of objects possessing feature k.



Defining a prior

> Define a prior on π by assuming that each π_k follows a Beta distribution (conjugate to the binomial):

$$p(\pi_k) = \text{Beta}(\pi_k; r, s) = \frac{\pi_k^{r-1} (1 - \pi_k)^{s-1}}{B(r, s)}$$

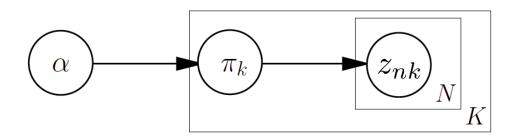
where B(r,s) is the beta function

$$B(r,s) = \int_0^1 \pi_k^{r-1} (1 - \pi_k)^{s-1} d\pi_k = \frac{\Gamma(r)\Gamma(s)}{\Gamma(r+s)}$$

We set r=lpha/K and s=1, so this equation becomes

$$B(\frac{\alpha}{K}, 1) = \frac{\Gamma(\frac{\alpha}{K})}{\Gamma(1 + \frac{\alpha}{K})} = \frac{K}{\alpha}$$





- Resulting probability model
 - Finite Beta-Bernoulli model

$$\pi_k | \alpha \sim \text{Beta}(\frac{\alpha}{K}, 1)$$
 $z_{nk} | \pi_k \sim \text{Bernoulli}(\pi_k)$

> Each z_{nk} is independent of all other assignments conditioned on π_k and the π_k are generated independently.



- We can now marginalize out π
 - Marginal probability of the matrix Z:

$$p(\mathbf{Z}) = \prod_{k=1}^{K} \int \left(\prod_{n=1}^{N} p(z_{nk} | \pi_k) \right) p(\pi_k) d\pi_k$$

$$= \prod_{k=1}^{K} \frac{B(m_k + \frac{\alpha}{K}, N - m_k + 1)}{B(\frac{\alpha}{K}, 1)}$$

$$= \prod_{k=1}^{K} \frac{\frac{\alpha}{K} \Gamma(m_k + \frac{\alpha}{K}) \Gamma(N - m_k + 1)}{\Gamma(N + 1 + \frac{\alpha}{K})}$$

conjugacy b/w binomial and beta

- \Rightarrow This distribution depends only on the counts m_k .
- ⇒ It is therefore exchangeable.



Important Property

- Bound on the number of entries
 - Expectation of the number of non-zero entries in Z:

$$\mathbb{E}\left[\mathbf{1}^T\mathbf{z}_k\mathbf{1}
ight] \ = \ \mathbb{E}\left[\sum_{n=1}^N\sum_{k=1}^Kz_{nk}
ight] = K\mathbb{E}\left[\mathbf{1}^T\mathbf{z}_k
ight] \qquad egin{array}{l} ext{(columns of } \ \mathbf{Z} ext{ are independent)} \ & ext{dependent)} \ & ext{dependent} \ & ext{depend$$

 $= K \sum_{n=1}^{N} \mathbb{E}[z_{nk}] = K \sum_{n=1}^{N} \int_{0}^{1} \pi_{k} p(\pi_{k}) d\pi_{k}$

Expectation of a $\mathrm{Beta}(r,s)$ random variable is r/(r+s)

$$= KN \frac{\frac{\alpha}{K}}{1 + \frac{\alpha}{K}} = \frac{N\alpha}{1 + \alpha/K}$$

 \Rightarrow For any K, the expectation of this number is bounded by $N\alpha$.

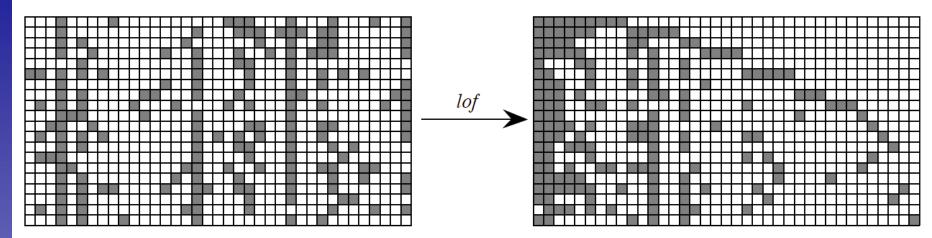


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Equivalence Class of Binary Matrices



Equivalence class of binary matrices

- > Define a function $\log(\mathbf{Z})$ that maps binary matrices into left-ordered binary matrices by ordering the columns of \mathbf{Z} by the magnitude of the binary number expressed by that column.
- There is a unique left-ordered form for every binary matrix.
- lacksquare Two matrices ${f Y}$ and ${f Z}$ are equivalent iff $\log({f Y})=\log({f Z})$.
- imes The \log -equivalence class of ${f Z}$ is denoted $[{f Z}]$.



Equivalence Class of Binary Matrices

- What is the cardinality of [Z]?
 - Columns of a binary matrix are not guaranteed to be unique:
 - Since an object can possess multiple features, it is possible for two features to be possessed by exactly the same set of objects.
 - > The cardinality of $[\mathbf{Z}]$ is therefore reduced if \mathbf{Z} contains identical columns

$$\begin{pmatrix} K \\ K_0, ..., K_{2^N - 1} \end{pmatrix} = \frac{K!}{\prod_{h=0}^{2^N - 1} K_h!}$$

where K_h is the number of columns with binary number h.



Towards Infinite Feature Models

- Taking the limit $K \to \infty$
 - > Probability of a lof-equivalence class of binary matrices

$$p([\mathbf{Z}]) = \sum_{\mathbf{Z} \in [\mathbf{Z}]} p(\mathbf{Z}) = \frac{K!}{\prod_{h=0}^{2^N - 1} K_h!} \prod_{k=1}^K \frac{\frac{\alpha}{K} \Gamma(m_k + \frac{\alpha}{K}) \Gamma(N - m_k + 1)}{\Gamma(N + 1 + \frac{\alpha}{K})}$$

Reordering the columns such that $m_k>0$ if $k\leq K_+$, we can derive (after several intermediate steps)

$$\lim_{K \to \infty} p([\mathbf{Z}]) = \frac{\alpha^{K_+}}{\prod_{h=0}^{2^N - 1} K_h!} \exp\{-\alpha H_N\} \prod_{k=1}^{K_+} \frac{(N - m_k)!(m_k - 1)!}{N!}$$

where H_N is the N^{th} harmonic number H_N = $\sum_{j=1}^N 1/j$.



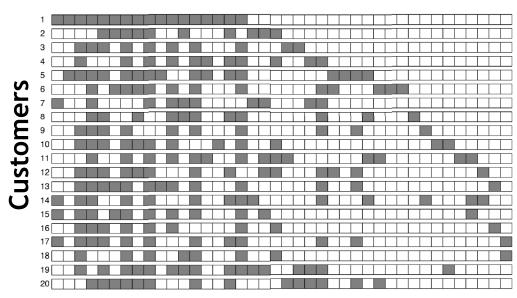
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The Indian Buffet Process



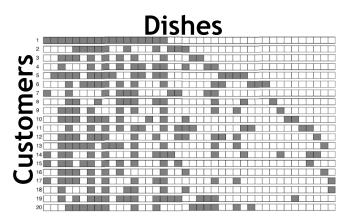


"Many Indian restaurants in London offer lunchtime buffets with an apparently infinite number of dishes"

[Zoubin Ghahramani]



The Indian Buffet Process



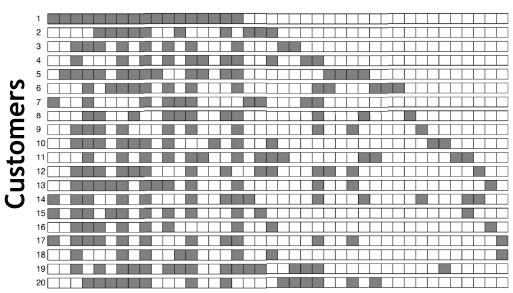
Analogy to Chinese Restaurant Process

- Visualize feature assignment as a sequential process of customers sampling dishes from an (infinitely long) buffet
- 1st customer starts at the left of the buffet, and takes a serving from each dish, stopping after a Poisson() number of dishes as her plate becomes overburdened.
- The $n^{\rm th}$ customer moves along the buffet, sampling dishes in proportion to their popularity, serving himself with probability m_k/n , and trying a ${
 m Poisson}(\alpha/n)$ number of new dishes.
- \succ The customer-dish matrix is our feature matrix, ${f Z}_{f \cdot}$



The Indian Buffet Process (IBP)





Properties of the IBP

- Generative process to create samples from an infinite latent feature model.
- The IBP is exchangeable, up to a permutation of the order with which dishes are listed in the feature matrix.
- lacksquare The number of features sampled at least once is $\mathcal{O}(lpha \, \log \, N)$.



to be continued in 2013



References and Further Reading

- Tutorial papers for infinite latent factor models
 - A good introduction to the topic
 - Z. Ghahramani, T.L. Griffiths, P. Sollich, "<u>Bayesian Nonparametric</u> <u>Latent Feature Models</u>", Bayesian Statistics, 2006.