

Advanced Machine Learning Lecture 15

Latent Factor Models

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This Lecture: Advanced Machine Learning

- Regression Approaches
 - Linear Regression
 - Regularization (Ridge, Lasso)
 - Kernels (Kernel Ridge Regression)
 - Gaussian Processes
- Bayesian Estimation & Bayesian Non-Parametrics
 - > Prob. Distributions, Approx. Inference
 - » Mixture Models & EM
 - Dirichlet Processes
 - > Latent Factor Models
 - Beta Processes

Learning Winter'12

Advanced Machine

- SVMs and Structured Output Learning
 - SV Regression, SVDD
 - Large-margin Learning







 $f: \mathcal{X} \to \mathcal{Y}$



Topics of This Lecture

Latent Factor Models

- Motivation
- Example: PCA
- > Applications of PCA
- > Probabilistic PCA
- Maximum Likelihood for PCA
- > Other Latent Factor Models: FA, ICA

Towards Infinite Latent Factor Models

- General formulation
- Sparse latent factor models
- > Priors on binary matrices
- Finite latent feature model

Mixture Models vs. Latent Factor Models

- Mixture Models
 - > Assume that each observation was generated by exactly one of $K\ {\rm components.}$
 - > The uncertainty is just about which component is responsible.
- Latent Factor Models
 - Weaken this assumption.
 - Each observation is influenced by each of K components (factors or features) in a different way.
 - Sparse factor models: only a small subset of factors is active for each observation.



Latent Factor/Feature Models

- Most popular examples
 - Principal Component Analysis (PCA)
 - Factor Analysis (FA)
 - Independent Component Anlalysis (ICA)
- Properties
 - > All of those assume that the number of factors K is known.
 - > Usually, K is smaller than the dimensionality of the data: $K \ll D$
 - \Rightarrow Models provide dimensionality reduction.
- Let's look at PCA and see how it fits into this framework...



• Goal

- Solution Given a data set $X = \{\mathbf{x}_n\}$ in D dimensions, find the K-dimensional projection (K < D) that maximizes the variance of the projected data.
- Intuition: preserve as much variance as possible.
- One-dimensional example
 - > Project each data point \mathbf{x}_n onto the unit vector \mathbf{u}_1

$$y_n = \mathbf{u}_1^T \mathbf{x}_n$$

> What is the vector \mathbf{u}_1 that maximizes the variance of the projected data?



- One-dimensional example (cont'd)
 - > Mean of the projected data

$$\bar{y} = \mathbf{u}_1^T \bar{\mathbf{x}} \qquad \bar{\mathbf{x}} = \frac{1}{N} \sum_{n=1}^N \mathbf{x}_n$$

 ΛT

Variance of the projected data

$$\frac{1}{N} \sum_{n=1}^{N} \{\mathbf{u}_{1}^{T} \mathbf{x}_{n} - \mathbf{u}_{1}^{T} \bar{\mathbf{x}}\}^{2} = \mathbf{u}_{1}^{T} \mathbf{S} \mathbf{u}_{1}^{T}$$
$$\mathbf{S} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_{n} - \bar{\mathbf{x}}) (\mathbf{x}_{n} - \bar{\mathbf{x}})^{T}$$

where ${\bf S}$ is the data covariance matrix.



- Optimization problem
 - > Maximize the projected variance $\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1$ w.r.t. \mathbf{u}_1 .
 - \succ Problem: trivial solution is $\|\mathbf{u}_1\|
 ightarrow \infty$.
 - \Rightarrow Need to enforce the normalization condition $\mathbf{u}_1^{\mathrm{T}}\mathbf{u}_1 = 1$.
 - Formulation with Lagrange multiplier

$$\arg \max_{\mathbf{u}_1} \quad \mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 + \lambda_1 (1 - \mathbf{u}_1^T \mathbf{u}_1)$$

Setting the derivative to zero

$$\mathbf{Su}_1 = \lambda_1 \mathbf{u}_1$$

 \Rightarrow Eigenvalue problem: \mathbf{u}_1 must be eigenvector of \mathbf{S} .

$$\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 = \mathbf{u}_1^T \lambda_1 \mathbf{u}_1 = \lambda_1 \mathbf{u}_1^T \mathbf{u}_1 = \lambda_1$$

 \Rightarrow Maximal variance if λ_1 is the largest eigenvalue of S.



- General case
 - Inductively, we can show that the optimal linear projection into a K-dimensional space is given by the first K eigenvectors u₁,...,u_K of S.

$$\mathbf{y}_n = \mathbf{U}_{1..K} \mathbf{x}_n$$

Graphical interpretation





Uses of PCA

- Dimensionality reduction
 - > Work in a subspace that contains only the K most important dimensions.
 - > Advantages: faster processing, reduced memory footprint, robustness to noise.

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Example: Eigenfaces



Slide credit: Peter Belhumeur



Uses of PCA

- Dimensionality reduction
 - > Work in a subspace that contains only the K most important dimensions.
 - > Advantages: faster processing, reduced memory footprint, robustness to noise.

Data Preprocessing

- Remove correlations between different dimensions of the data and bring them to a common scale.
- Many classification or regression algorithms work better when the data is standardized, i.e., when each variable has zero mean and unit variance.
- Using PCA, we can make a more substantial normalization of the data to give it zero mean and unit covariance. This is known as whitening.



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PCA for Whitening

- Whitening procedure
 - Rewrite the eigenvector equation in matrix form

$$\mathbf{Su}_1 = \lambda_1 \mathbf{u}_1 \qquad \Rightarrow \qquad \mathbf{SU} = \mathbf{UL}$$

where $\mathbf{L} = \operatorname{diag}\{\lambda_i\}, \ \mathbf{U} = [\mathbf{u}_1, ..., \mathbf{u}_D].$

Define for each data point the transformed value as

$$\mathbf{y}_n = \mathbf{L}^{-1/2} \mathbf{U}^T (\mathbf{x}_n - \bar{\mathbf{x}})$$

 \Rightarrow The transformed set $\{\mathbf{y}_n\}$ has zero mean and unit covariance.

$$\frac{1}{N} \sum_{n=1}^{N} \mathbf{y}_n \mathbf{y}_n^T = \frac{1}{N} \sum_{n=1}^{N} \mathbf{L}^{-1/2} \mathbf{U}^T (\mathbf{x} - \bar{\mathbf{x}}) (\mathbf{x} - \bar{\mathbf{x}})^T \mathbf{U} \mathbf{L}^{-1/2}$$
$$= \mathbf{L}^{-1/2} \mathbf{U}^T \mathbf{S} \mathbf{U} \mathbf{L}^{-1/2} = \mathbf{L}^{-1/2} \mathbf{L} \mathbf{L}^{-1/2} = \mathbf{I}$$



Whitening Example



- Whitening result
 - Correlations are removed.
 - Distances are normalized to same value range.



- Discussion
 - The formulation of PCA we have just seen was based on a linear projection of data into a lower-dim. subspace.
 - We now show that PCA can also be expressed as the ML solution of a probabilistic latent variable model.
- Advantages of Probabilistic PCA
 - We can derive an EM algorithm that is efficient in situations where only few leading eigenvectors are required.
 - Probabilistic model + EM makes it possible to deal with missing data values.
 - Basis for a Bayesian treatment of PCA in which the dimensionality of the principal subspace can be found automatically.



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Probabilistic PCA

- Graphical Model
 - Introduce an explicit latent variable z corresponding to the principal component subspace.
 - Define a Gaussian prior distribution

 $p(\mathbf{z}) = \mathcal{N}(\mathbf{z}|\mathbf{0},\mathbf{I})$

Conditional distribution also Gaussian

$$p(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \sigma^2 \mathbf{I})$$



⇒ Example of a Linear Gaussian framework: all of the marginal and conditional distributions are Gaussian

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As we will see, the columns of W span an K-dimensional linear subspace within the data space that corresponds to the principal subspace.



- Generative interpretation
 - D-Dimensional observed variable x is defined by a linear transformation of the K-dimensional latent variable z, plus some added (isotropic Gaussian) noise.

$$\mathbf{x} = \mathbf{W}\mathbf{z} + \boldsymbol{\mu} + \boldsymbol{\epsilon}$$

Marginal distribution

$$p(\mathbf{x}) = \int p(\mathbf{x}|\mathbf{z})p(\mathbf{z})\mathrm{d}\mathbf{z}$$

> Because of the linear-Gaussian model, this will again be Gaussian $p({\bf x}) = \mathcal{N}({\bf x}|{\bm \mu}, {\bf C})$

where $\mathbf{C} = \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I}$



- Properties
 - > There is a rotational ambiguity in the parametrization.
 - > Consider a rotation of the latent parameter space with orthonormal matrix ${f R}$ (orthogonality property: ${f R}{f R}^{
 m T}={f I}$).

$$\widetilde{\mathbf{W}} = \mathbf{W}\mathbf{R}$$

$$\widetilde{\mathbf{W}}\widetilde{\mathbf{W}}^T = \mathbf{W}\mathbf{R}\mathbf{R}^T\mathbf{W}^T = \mathbf{W}\mathbf{W}^T$$

- \Rightarrow Thus, the covariance matrix ${\bf C}$ is independent of ${\bf R}.$
- > Efficiency trick: instead of evaluating ${f C}^{-1}$ directly, use the following equivalence (${\cal O}(D^3) o {\cal O}(K^3)$).

$$\mathbf{C}^{-1} = \sigma^{-2}\mathbf{I} - \sigma^{-2}\mathbf{W}\mathbf{M}^{-1}\mathbf{W}^{T}$$

with

$$\mathbf{M} = \mathbf{W}^T \mathbf{W} + \sigma^2 \mathbf{I}$$



- Posterior distribution
 - > Can again be derived from properties of linear Gaussian models

$$p(\mathbf{z}|\mathbf{x}) = \mathcal{N}\left(\mathbf{z}|\mathbf{M}^{-1}\mathbf{W}^{T}(\mathbf{x}-\boldsymbol{\mu}), \sigma^{2}\mathbf{M}^{-1}\right)$$



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Maximum Likelihood for PCA

- Maximum Likelihood estimate
 - Log-likelihood function

$$\log p(\mathbf{X}|\mathbf{W}, \boldsymbol{\mu}, \sigma^2) = \sum_{n=1}^N \log p(\mathbf{x}_n | \mathbf{W}, \boldsymbol{\mu}, \sigma^2)$$
$$= -\frac{1}{2} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu})^T \mathbf{C}^{-1} (\mathbf{x}_n - \boldsymbol{\mu})$$
$$-\frac{ND}{2} \log(2\pi) - \frac{N}{2} \log |\mathbf{C}|$$

> Optimizing the parameters

$$\frac{\partial}{\partial \boldsymbol{\mu}} p(\mathbf{X} | \mathbf{W}, \boldsymbol{\mu}, \sigma^2) \stackrel{!}{=} 0 \qquad \Rightarrow \qquad \boldsymbol{\mu} = \bar{\mathbf{x}}$$



Maximum Likelihood for PCA

- Maximum Likelihood estimate
 - > Plugging in the result for μ ...

$$\log p(\mathbf{X}|\boldsymbol{\mu}, \mathbf{W}, \sigma^2) = -\frac{N}{2} \left\{ D \log(2\pi) + \log |\mathbf{C}| + \operatorname{Tr}(\mathbf{C}^{-1}\mathbf{S}) \right\}$$

Maximizing w.r.t. W yields a closed-form solution:

$$\mathbf{W}_{\mathrm{ML}} = \mathbf{U}_K (\mathbf{L}_K - \sigma^2 \mathbf{I})^{1/2} \mathbf{R}$$

> where

- \mathbf{U}_K is a $D \times K$ matrix, whose columns are given by the K principal eigenvectors of the data covariance matrix \mathbf{S} ,
- L contains eigenvalues λ_i , and
- **R** is an arbitrary $K \times K$ rotation matrix.
- $\Rightarrow \text{The columns of W define the principal subspace of standard} \\ \text{PCA. For } \mathbf{R} = \mathbf{I}, \text{ they correspond to the principal eigenvectors} \\ [\mathbf{u}_1, ..., \mathbf{u}_K], \text{ scaled by the variance parameters } \lambda_i \sigma^2. \end{cases}$



Maximum Likelihood for PCA

- Maximum Likelihood estimate (cont'd)
 - > Maximizing w.r.t. σ :

$$\sigma_{\rm ML}^2 = \frac{1}{D-K} \sum_{i=K+1}^{D} \lambda_i$$

 $\Rightarrow \sigma^2{}_{\rm ML}$ is the average variance associated with the discarded dimensions.

Interpretation of Probabilistic PCA

- Putting all those results together...
 - Consider again the covariance matrix

$$\mathbf{C} = \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I}$$

where

$$\mathbf{W}_{\mathrm{ML}} = \mathbf{U}_{K} (\mathbf{L}_{K} - \sigma^{2} \mathbf{I})^{1/2} \mathbf{R}$$
$$\sigma_{\mathrm{ML}}^{2} = \frac{1}{D - K} \sum_{i=K+1}^{D} \lambda_{i}$$

- \Rightarrow The model correctly captures the variance of the data along the principal axes and approximates the variance in all remaining directions by σ^2 , the average of the discarded eigenvalues.
- To construct C, we simply set R = I and compute the principal eigenvalues and eigenvectors of the data covariance matrix S.
 If C is obtained in a different way. R may still be arbitrary.
- \Rightarrow If ${\bf C}$ is obtained in a different way, ${\bf R}$ may still be arbitrary.

Discussion: PCA vs. Probabilistic PCA

- Comparison with standard PCA:
 - > PCA is generally formulated as a projection of points from the D-dimensional space onto a K-dimensional linear subspace.
 - Probabilistic PCA is more naturally expressed as a mapping from the latent space into the data space via

$$\mathbf{x} = \mathbf{W}\mathbf{z} + \boldsymbol{\mu} + \boldsymbol{\epsilon}$$

For applications such as visualization or data compression, we can reverse this mapping using Bayes' theorem.

 $\mathbb{E}[\mathbf{z}|\mathbf{x}] = \mathbf{M}^{-1} \mathbf{W}_{\mathrm{ML}}^T (\mathbf{x} - \bar{\mathbf{x}}) \quad \text{ where } \quad \mathbf{M} = \mathbf{W}^T \mathbf{W} + \sigma^2 \mathbf{I}$

This projects to a point in data space given by

$$\mathbf{W}\mathbb{E}[\mathbf{z}|\mathbf{x}] + \boldsymbol{\mu}$$

 \succ In the limit $\sigma
ightarrow 0$, this reduces to the standard PCA model.



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Other Latent Factor Models

- Factor Analysis (FA)
 - Linear-Gaussian latent variable model, closely related to Probabilistic PCA.
 - Probabilistic PCA uses an isotropic covariance

 $p(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \sigma^2 \mathbf{I})$

Factor Analysis instead assumes a diagonal covariance

$$p(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \boldsymbol{\Psi}) \qquad \boldsymbol{\Psi} = \operatorname{diag}\{\psi_i\}$$

- > The FA model explains the observed covariance structure of the data by representing the independent variables associated with each coordinate by the matrix Ψ and capturing the covariance between variables in the matrix W.
- > In the literature, the columns of W are called factor loadings, the diagonal elements ψ_i are called uniquenesses.



Other Latent Factor Models (2)

- Independent Component Analysis (ICA)
 - Model for which the observed variables are related linearly to the latent variables, but for which the latent distribution is non-Gaussian.
 - > Consider a distribution over latent variables that factorizes

$$p(\mathbf{z}) = \prod_{j=1}^{K} p(z_j)$$

- i.e., the components z_j are independent.
- > This definition requires that the latent variables have a non-Gaussian distribution (as Gaussian models always have the rotational ambiguity R in latent space).
- There is a large variety of ICA models and corresponding algorithms, differing mainly in the choice of latent-variable distribution.

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Next Steps from Here...

Discussion

- > We have now derived that the PCA result can be obtained as the ML estimate of the corresponding probabilistic model.
- This result can directly be used to incorporate priors and derive a Bayesian extension of the model.
- > We can do similar things for FA and ICA...
- In the following, we will go into a different direction
 - » What happens when we let $K \to \infty$?
 - > Can we automatically determine K?

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General Latent Factor Models

- General formulation
 - Assume that the data are generated by noisy weighted combination of latent factors

$$\mathbf{x}_n = \mathbf{F}\mathbf{y}_n + oldsymbol{\epsilon}$$

- > E.g., in Factor Analysis, F would be a $D \times K$ factor loading matrix expressing how latent factor k influences observation dimension d. y_n would be a K-dimensional vector expressing the activity of each factor.
- Advantages of latent feature modeling
 - > Each group of observations is associated with a subset of the possible latent features/factors.
 - Factorial power: There are 2^K combinations of K features, while accurate mixture modeling may require many more clusters.



Sparse Latent Factor Models

- Goal: Infinite models
 - > We would like to work with infinite-dimensional models ($K{
 ightarrow}\infty$)
 - > In order to do keep inference tractable, however, we have to restrict the model somehow.
 - Mixture Models: DPs enforce that the main part of the probability mass is concentrated on few cluster components.
 - Latent Factor Models: enforce that each object is represented by a sparse subset of an unbounded number of features.
 - Incorporating sparsity
 - > Decompose \mathbf{F} into the product of two components: $\mathbf{F} = \mathbf{Z} \otimes \mathbf{W}$, where \otimes is the Hadamard product (element-wise product).
 - z_{mk} is a binary mask variable indicating whether factor k is "on".
 - w_{mk} is a continuous weight variable.
 - \Rightarrow Enforce sparsity by restricting the non-zero entries in ${f Z}.$



Sparse Latent Factor Models



- Latent feature modeling
 - In PCA (FA, ICA, etc.), objects have non-zero values on every feature and every entry of Z is 1.
 - > In sparse latent feature models, only a sparse subset of features take non-zero values, and Z makes those subsets explicit.

Towards a Full Bayesian Treatment

- Inference in Latent Feature Models
 - Goal: Infer the latent factors, mask variables, and weights.
 - Classical approaches (PCA, FA, ICA) fit point estimates of the parameters through ML estimation.
- Bayesian approach
 - > Specify a prior over latent features/factors $p(\mathbf{F})$ and a distribution over observed property distributions $p(\mathbf{X}|\mathbf{F})$.
 - > Compute the posterior $\, p({f F},{f Z},{f W}|{f X}) \! . \,$
 - > Our focus will be on $p(\mathbf{F}) = p(\mathbf{Z})p(\mathbf{W})$, showing how such a prior can be defined without placing an upper bound on the number of features/factors.



Priors on Binary Matrices

- Let's first go back to DPs/CRPs
 - Back there, we also had binary \geq matrices due to 1-of-K coding.
 - What is different here? \succ
 - **Binary matrices for clustering**
 - We can think of CRPs as priors on infinite binary matrices, where...
 - …each data point is assigned to one and only one cluster (class).
 - ...the rows sum to one.



Clusters Image source: Zoubin Gharahmani

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Priors on Binary Matrices

- Let's first go back to DPs/CRPs
 - > Back there, we also had binary matrices due to 1-of-K coding.
 - What is different here?
- More general binary matrices
 - Each data point can now have multiple factors/features.
 - The rows sum to more than one.
 - ⇒ What is the corresponding prior on infinite binary matrices?



Factors/Features

Slide adapted from Zoubin Gharahmani

Image source: Zoubin Gharahmani

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Priors on Latent Factor Models

- Defining suitable priors
 - > We will focus on defining a prior on Z, since the effective dimensionality of the latent feature model is determined by Z.
 - > Assuming that Z is sparse, we can define a prior for infinite latent feature models by defining a distribution over infinite binary matrices.
- Desiderata for such a distribution
 - > Objects should be exchangeable.
 - Inference should be tractable.

• Procedure

- Start with a model that assumes a finite number of features and consider the limit as this number approaches infinity.
- \Rightarrow Next lecture...



References and Further Reading

• More information on latent factor models and particularly PCA can be found in Chapter 12 of

Christopher M. Bishop Pattern Recognition and Machine Learning Springer, 2006



- Tutorial papers for infinite latent factor models
 - A good introduction to the topic
 - Z. Ghahramani, T.L. Griffiths, P. Sollich, "<u>Bayesian Nonparametric</u> <u>Latent Feature Models</u>", Bayesian Statistics, 2006.