

RWTH AACHEN  
UNIVERSITY

# Machine Learning - Lecture 2

## Probability Density Estimation

15.04.2016

Bastian Leibe  
RWTH Aachen  
<http://www.vision.rwth-aachen.de>  
leibe@vision.rwth-aachen.de

Many slides adapted from B. Schiele

Machine Learning Summer '16

RWTH AACHEN  
UNIVERSITY

## Announcements

- Course webpage
  - <http://www.vision.rwth-aachen.de/teaching/>
  - Slides will be made available on the webpage
- L2P electronic repository
  - Exercises and supplementary materials will be posted on the L2P
- Please subscribe to the lecture on the Campus system!
  - Important to get email announcements and L2P access!

B. Leibe

2

Machine Learning Summer '16

RWTH AACHEN  
UNIVERSITY

## Course Outline

- Fundamentals (2 weeks)
  - Bayes Decision Theory
  - Probability Density Estimation
- Discriminative Approaches (5 weeks)
  - Linear Discriminant Functions
  - Support Vector Machines
  - Ensemble Methods & Boosting
  - Randomized Trees, Forests & Ferns
- Generative Models (4 weeks)
  - Bayesian Networks
  - Markov Random Fields

B. Leibe

4

Machine Learning Summer '16

RWTH AACHEN  
UNIVERSITY

## Topics of This Lecture

- Recap: Bayes Decision Theory
  - Basic concepts
  - Minimizing the misclassification rate
  - Minimizing the expected loss
  - Discriminant functions
- Probability Density Estimation
  - General concepts
  - Gaussian distribution
- Parametric Methods
  - Maximum Likelihood approach
  - Bayesian vs. Frequentist views on probability
  - Bayesian Learning

B. Leibe

5

Machine Learning Summer '16

RWTH AACHEN  
UNIVERSITY

## Recap: Bayes Decision Theory Concepts

- Concept 1: Priors (a priori probabilities)  $p(C_k)$ 
  - What we can tell about the probability *before seeing the data*.
  - Example:
 

a a b a b a a b a  
 b a a a a b a a b a  
 a b a a a b b a  
 b a b a a b a a

$P(a)=0.75$   
 $P(b)=0.25$

?

$C_1 = a \quad p(C_1) = 0.75$   
 $C_2 = b \quad p(C_2) = 0.25$
- In general:  $\sum_k p(C_k) = 1$

B. Leibe

6

Machine Learning Summer '16

RWTH AACHEN  
UNIVERSITY

## Recap: Bayes Decision Theory Concepts

- Concept 2: Conditional probabilities  $p(x | C_k)$ 
  - Let  $x$  be a feature vector.
  - $x$  measures/describes certain properties of the input.
    - E.g. number of black pixels, aspect ratio, ...
  - $p(x|C_k)$  describes its **likelihood** for class  $C_k$ .

$p(x|a)$

$x$

$p(x|b)$

$x$

B. Leibe

7

Machine Learning Summer '16

RWTH AACHEN UNIVERSITY

## Bayes Decision Theory Concepts

- **Concept 3: Posterior probabilities**  $p(C_k | x)$ 
  - We are typically interested in the *a posteriori* probability, i.e. the probability of class  $C_k$  given the measurement vector  $x$ .
- **Bayes' Theorem:**

$$p(C_k | x) = \frac{p(x | C_k) p(C_k)}{p(x)} = \frac{p(x | C_k) p(C_k)}{\sum_i p(x | C_i) p(C_i)}$$
- **Interpretation**

$$\text{Posterior} = \frac{\text{Likelihood} \times \text{Prior}}{\text{Normalization Factor}}$$

Machine Learning Summer'16 8

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Recap: Bayes Decision Theory

Machine Learning Summer'16 9

Slide credit: Bernt Schiele B. Leibe Image source: C.M. Bishop, 2009

RWTH AACHEN UNIVERSITY

## Recap: Bayes Decision Theory

- **Optimal decision rule**
  - Decide for  $C_1$  if
 
$$p(C_1 | x) > p(C_2 | x)$$
  - This is equivalent to
 
$$p(x | C_1) p(C_1) > p(x | C_2) p(C_2)$$
  - Which is again equivalent to (**Likelihood-Ratio test**)
 
$$\frac{p(x | C_1)}{p(x | C_2)} > \underbrace{\frac{p(C_2)}{p(C_1)}}_{\text{Decision threshold } \theta}$$

Machine Learning Summer'16 10

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Bayes Decision Theory

- **Decision regions:  $\mathcal{R}_1, \mathcal{R}_2, \mathcal{R}_3, \dots$**

Machine Learning Summer'16 11

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Recap: Minimizing the Expected Loss

- **Example:**
  - 2 Classes:  $C_1, C_2$
  - 2 Decision:  $\alpha_1, \alpha_2$
  - Loss function:  $L(\alpha_j | C_k) = L_{kj}$
  - Expected loss (= risk  $R$ ) for the two decisions:
 
$$\mathbb{E}_{\alpha_1}[L] = R(\alpha_1 | \mathbf{x}) = L_{11}p(C_1 | \mathbf{x}) + L_{21}p(C_2 | \mathbf{x})$$

$$\mathbb{E}_{\alpha_2}[L] = R(\alpha_2 | \mathbf{x}) = L_{12}p(C_1 | \mathbf{x}) + L_{22}p(C_2 | \mathbf{x})$$
- **Goal: Decide such that expected loss is minimized**
  - I.e. decide  $\alpha_1$  if  $R(\alpha_2 | \mathbf{x}) > R(\alpha_1 | \mathbf{x})$

Machine Learning Summer'16 12

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Recap: Minimizing the Expected Loss

$$R(\alpha_2 | \mathbf{x}) > R(\alpha_1 | \mathbf{x})$$

$$L_{12}p(C_1 | \mathbf{x}) + L_{22}p(C_2 | \mathbf{x}) > L_{11}p(C_1 | \mathbf{x}) + L_{21}p(C_2 | \mathbf{x})$$

$$(L_{12} - L_{11})p(C_1 | \mathbf{x}) > (L_{21} - L_{22})p(C_2 | \mathbf{x})$$

$$\frac{(L_{12} - L_{11})}{(L_{21} - L_{22})} > \frac{p(C_2 | \mathbf{x})}{p(C_1 | \mathbf{x})} = \frac{p(\mathbf{x} | C_2) p(C_2)}{p(\mathbf{x} | C_1) p(C_1)}$$

$$\frac{p(\mathbf{x} | C_1)}{p(\mathbf{x} | C_2)} > \frac{(L_{21} - L_{22}) p(C_2)}{(L_{12} - L_{11}) p(C_1)}$$

⇒ Adapted decision rule taking into account the loss.

Machine Learning Summer'16 13

Slide credit: Bernt Schiele B. Leibe

Machine Learning Summer'16

## The Reject Option

- Classification errors arise from regions where the largest posterior probability  $p(\mathcal{C}_k|x)$  is significantly less than 1.
  - These are the regions where we are relatively uncertain about class membership.
  - For some applications, it may be better to reject the automatic decision entirely in such a case and e.g. consult a human expert.

B. Leibe 14  
Image source: C.M. Bishop, 2009

Machine Learning Summer'16

## Discriminant Functions

- Formulate classification in terms of comparisons
  - Discriminant functions
 
$$y_1(x), \dots, y_K(x)$$
  - Classify  $x$  as class  $C_k$  if
 
$$y_k(x) > y_j(x) \quad \forall j \neq k$$
- Examples (Bayes Decision Theory)
 
$$y_k(x) = p(\mathcal{C}_k|x)$$

$$y_k(x) = p(x|\mathcal{C}_k)p(\mathcal{C}_k)$$

$$y_k(x) = \log p(x|\mathcal{C}_k) + \log p(\mathcal{C}_k)$$

Slide credit: Bernt Schiele B. Leibe  
15

Machine Learning Summer'16

## Different Views on the Decision Problem

- $y_k(x) \propto p(x|\mathcal{C}_k)p(\mathcal{C}_k)$ 
  - First determine the class-conditional densities for each class individually and separately infer the prior class probabilities.
  - Then use Bayes' theorem to determine class membership.
  - $\Rightarrow$  *Generative methods*
- $y_k(x) = p(\mathcal{C}_k|x)$ 
  - First solve the inference problem of determining the posterior class probabilities.
  - Then use decision theory to assign each new  $x$  to its class.
  - $\Rightarrow$  *Discriminative methods*
- Alternative
  - Directly find a discriminant function  $y_k(x)$  which maps each input  $x$  directly onto a class label.

B. Leibe 16

Machine Learning Summer'16

## Topics of This Lecture

- Bayes Decision Theory
  - Basic concepts
  - Minimizing the misclassification rate
  - Minimizing the expected loss
  - Discriminant functions
- Probability Density Estimation
  - General concepts
  - Gaussian distribution
- Parametric Methods
  - Maximum Likelihood approach
  - Bayesian vs. Frequentist views on probability
  - Bayesian Learning

B. Leibe 17

Machine Learning Summer'16

## Probability Density Estimation

- Up to now
  - Bayes optimal classification
  - Based on the probabilities  $p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)$
- How can we estimate (=learn) those probability densities?
  - Supervised training case: data and class labels are known.
  - Estimate the probability density for each class  $\mathcal{C}_k$  separately:
 
$$p(\mathbf{x}|\mathcal{C}_k)$$
  - (For simplicity of notation, we will drop the class label  $\mathcal{C}_k$  in the following.)

Slide credit: Bernt Schiele B. Leibe  
18

Machine Learning Summer'16

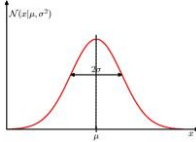
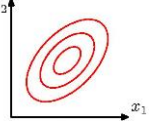
## Probability Density Estimation

- Data:  $x_1, x_2, x_3, x_4, \dots$
- Estimate:  $p(x)$
- Methods
  - Parametric representations (today)
  - Non-parametric representations (lecture 3)
  - Mixture models (lecture 4)

Slide credit: Bernt Schiele B. Leibe  
19

RWTH AACHEN UNIVERSITY

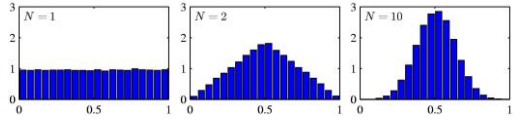
## The Gaussian (or Normal) Distribution

- One-dimensional case
  - Mean  $\mu$
  - Variance  $\sigma^2$
$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$$

- Multi-dimensional case
  - Mean  $\mu$
  - Covariance  $\Sigma$
$$\mathcal{N}(x|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2}|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right\}$$


20  
B. Leibe Image source: C.M. Bishop, 2006

RWTH AACHEN UNIVERSITY

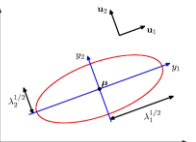
## Gaussian Distribution - Properties

- Central Limit Theorem
  - "The distribution of the sum of  $N$  i.i.d. random variables becomes increasingly Gaussian as  $N$  grows."
  - In practice, the convergence to a Gaussian can be very rapid.
  - This makes the Gaussian interesting for many applications.
- Example:  $N$  uniform  $[0, 1]$  random variables.
 

21  
B. Leibe Image source: C.M. Bishop, 2006

RWTH AACHEN UNIVERSITY

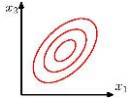
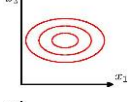
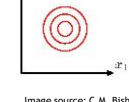
## Gaussian Distribution - Properties

- Quadratic Form
  - $\mathcal{N}$  depends on  $x$  through the exponent
  - $\Delta^2 = (x-\mu)^T \Sigma^{-1}(x-\mu)$
  - Here,  $\Delta$  is often called the Mahalanobis distance from  $x$  to  $\mu$ .
- Shape of the Gaussian
  - $\Sigma$  is a real, symmetric matrix.
  - We can therefore decompose it into its eigenvectors
  - $\Sigma = \sum_{i=1}^D \lambda_i u_i u_i^T$       $\Sigma^{-1} = \sum_{i=1}^D \frac{1}{\lambda_i} u_i u_i^T$
  - and thus obtain  $\Delta^2 = \sum_{i=1}^D \frac{y_i^2}{\lambda_i}$  with  $y_i = u_i^T(x-\mu)$ .
  - $\Rightarrow$  Constant density on ellipsoids with main directions along the eigenvectors  $u_i$  and scaling factors  $\sqrt{\lambda_i}$ .

22  
B. Leibe Image source: C.M. Bishop, 2006

RWTH AACHEN UNIVERSITY

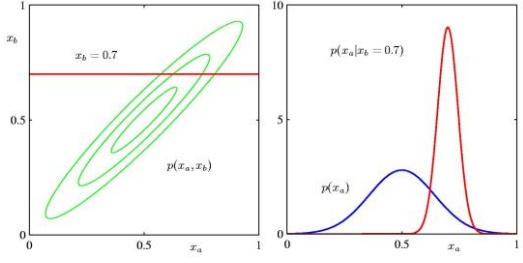
## Gaussian Distribution - Properties

- Special cases
  - Full covariance matrix  $\Sigma = [\sigma_{ij}]$ 
    - $\Rightarrow$  General ellipsoid shape
  - Diagonal covariance matrix  $\Sigma = \text{diag}\{\sigma_i\}$ 
    - $\Rightarrow$  Axis-aligned ellipsoid
  - Uniform variance  $\Sigma = \sigma^2 \mathbf{I}$ 
    - $\Rightarrow$  Hypersphere

23  
B. Leibe Image source: C.M. Bishop, 2006

RWTH AACHEN UNIVERSITY

## Gaussian Distribution - Properties

- The marginals of a Gaussian are again Gaussians:
 

24  
B. Leibe Image source: C.M. Bishop, 2006

RWTH AACHEN UNIVERSITY

## Topics of This Lecture

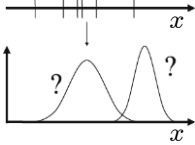
- Bayes Decision Theory
  - Basic concepts
  - Minimizing the misclassification rate
  - Minimizing the expected loss
  - Discriminant functions
- Probability Density Estimation
  - General concepts
  - Gaussian distribution
- Parametric Methods
  - Maximum Likelihood approach
  - Bayesian vs. Frequentist views on probability
  - Bayesian Learning

25  
B. Leibe

RWTH AACHEN UNIVERSITY

## Parametric Methods

- Given
  - Data  $X = \{x_1, x_2, \dots, x_N\}$
  - Parametric form of the distribution with parameters  $\theta$
  - E.g. for Gaussian distrib.:  $\theta = (\mu, \sigma)$
- Learning
  - Estimation of the parameters  $\theta$
- Likelihood of  $\theta$ 
  - Probability that the data  $X$  have indeed been generated from a probability density with parameters  $\theta$
$$L(\theta) = p(X|\theta)$$



26

Machine Learning Summer'16  
Slide adapted from Bernt Schiele  
B. Leibe

RWTH AACHEN UNIVERSITY

## Maximum Likelihood Approach

- Computation of the likelihood
  - Single data point:  $p(x_n|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x_n - \mu)^2}{2\sigma^2}\right\}$
  - Assumption: all data points are independent
$$L(\theta) = p(X|\theta) = \prod_{n=1}^N p(x_n|\theta)$$
- Log-likelihood
 
$$E(\theta) = -\ln L(\theta) = -\sum_{n=1}^N \ln p(x_n|\theta)$$
- Estimation of the parameters  $\theta$  (Learning)
  - Maximize the likelihood
  - Minimize the negative log-likelihood

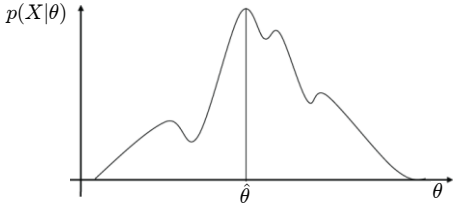
27

Machine Learning Summer'16  
Slide credit: Bernt Schiele  
B. Leibe

RWTH AACHEN UNIVERSITY

## Maximum Likelihood Approach

- Likelihood:  $L(\theta) = p(X|\theta) = \prod_{n=1}^N p(x_n|\theta)$
- We want to obtain  $\hat{\theta}$  such that  $L(\hat{\theta})$  is maximized.



28

Machine Learning Summer'16  
Slide credit: Bernt Schiele  
B. Leibe

RWTH AACHEN UNIVERSITY

## Maximum Likelihood Approach

- Minimizing the log-likelihood
  - How do we minimize a function?
    - Take the derivative and set it to zero.
$$\frac{\partial}{\partial \theta} E(\theta) = -\frac{\partial}{\partial \theta} \sum_{n=1}^N \ln p(x_n|\theta) = -\sum_{n=1}^N \frac{\frac{\partial}{\partial \theta} p(x_n|\theta)}{p(x_n|\theta)} \stackrel{!}{=} 0$$
- Log-likelihood for Normal distribution (1D case)
 
$$E(\theta) = -\sum_{n=1}^N \ln p(x_n|\mu, \sigma)$$

$$= -\sum_{n=1}^N \ln \left( \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{\|x_n - \mu\|^2}{2\sigma^2}\right\} \right)$$

29

Machine Learning Summer'16  
B. Leibe

RWTH AACHEN UNIVERSITY

## Maximum Likelihood Approach

- Minimizing the log-likelihood
 
$$\frac{\partial}{\partial \mu} E(\mu, \sigma) = -\sum_{n=1}^N \frac{\frac{\partial}{\partial \mu} p(x_n|\mu, \sigma)}{p(x_n|\mu, \sigma)}$$

$$= -\sum_{n=1}^N \frac{2(x_n - \mu)}{2\sigma^2}$$

$$= \frac{1}{\sigma^2} \sum_{n=1}^N (x_n - \mu)$$

$$= \frac{1}{\sigma^2} \left( \sum_{n=1}^N x_n - N\mu \right)$$

$$\frac{\partial}{\partial \mu} E(\mu, \sigma) \stackrel{!}{=} 0 \iff \hat{\mu} = \frac{1}{N} \sum_{n=1}^N x_n$$

$$p(x_n|\mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{\|x_n - \mu\|^2}{2\sigma^2}}$$

30

Machine Learning Summer'16  
B. Leibe

RWTH AACHEN UNIVERSITY

## Maximum Likelihood Approach

- We thus obtain
 
$$\hat{\mu} = \frac{1}{N} \sum_{n=1}^N x_n$$

"sample mean"
- In a similar fashion, we get
 
$$\hat{\sigma}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \hat{\mu})^2$$

"sample variance"
- $\hat{\theta} = (\hat{\mu}, \hat{\sigma})$  is the **Maximum Likelihood estimate** for the parameters of a Gaussian distribution.
- This is a very important result.
- Unfortunately, it is wrong...

31

Machine Learning Summer'16  
B. Leibe

RWTH AACHEN UNIVERSITY

## Maximum Likelihood Approach

- Or not wrong, but rather **biased**...
- Assume the samples  $x_1, x_2, \dots, x_N$  come from a true Gaussian distribution with mean  $\mu$  and variance  $\sigma^2$ 
  - We can now compute the expectations of the ML estimates with respect to the data set values. It can be shown that
 
$$\mathbb{E}(\mu_{ML}) = \mu$$

$$\mathbb{E}(\sigma_{ML}^2) = \left(\frac{N-1}{N}\right)\sigma^2$$
  - ⇒ The ML estimate will underestimate the true variance.
- Corrected estimate:
 
$$\tilde{\sigma}^2 = \frac{N}{N-1}\sigma_{ML}^2 = \frac{1}{N-1}\sum_{n=1}^N (x_n - \hat{\mu})^2$$

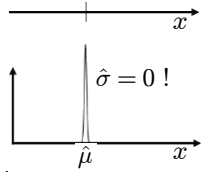
32

RWTH AACHEN UNIVERSITY

## Maximum Likelihood - Limitations

- Maximum Likelihood has several significant limitations
  - It systematically underestimates the variance of the distribution!
  - E.g. consider the case
 
$$N = 1, X = \{x_1\}$$

⇒ Maximum-likelihood estimate:




- We say ML *overfits to the observed data*.
- We will still often use ML, but it is important to know about this effect.

33

RWTH AACHEN UNIVERSITY

## Deeper Reason

- Maximum Likelihood is a **Frequentist** concept
  - In the **Frequentist view**, probabilities are the frequencies of random, repeatable events.
  - These frequencies are fixed, but can be estimated more precisely when more data is available.
- This is in contrast to the **Bayesian** interpretation
  - In the **Bayesian view**, probabilities quantify the uncertainty about certain states or events.
  - This uncertainty can be revised in the light of new evidence.
- Bayesians and Frequentists do not like each other too well...



34

RWTH AACHEN UNIVERSITY

## Bayesian vs. Frequentist View

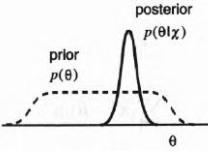
- To see the difference...
  - Suppose we want to estimate the uncertainty whether the Arctic ice cap will have disappeared by the end of the century.
  - This question makes no sense in a Frequentist view, since the event cannot be repeated numerous times.
  - In the Bayesian view, we generally have a prior, e.g. from calculations how fast the polar ice is melting.
  - If we now get fresh evidence, e.g. from a new satellite, we may revise our opinion and update the uncertainty from the prior.
 
$$\text{Posterior} \propto \text{Likelihood} \times \text{Prior}$$
  - This generally allows to get better uncertainty estimates for many situations.
- Main Frequentist criticism
  - The prior has to come from somewhere and if it is wrong, the result will be worse.

35

RWTH AACHEN UNIVERSITY

## Bayesian Approach to Parameter Learning

- Conceptual shift
  - Maximum Likelihood views the true parameter vector  $\theta$  to be unknown, but fixed.
  - In Bayesian learning, we consider  $\theta$  to be a random variable.
- This allows us to use knowledge about the parameters  $\theta$ 
  - i.e. to use a prior for  $\theta$
  - Training data then converts this prior distribution on  $\theta$  into a posterior probability density.



- The prior thus encodes knowledge we have about the type of distribution we expect to see for  $\theta$ .

36

RWTH AACHEN UNIVERSITY

## Bayesian Learning Approach

- Bayesian view:
  - Consider the parameter vector  $\theta$  as a random variable.
  - When estimating the parameters, what we compute is
 
$$p(x|X) = \int p(x, \theta|X) d\theta$$

Assumption: given  $\theta$ , this doesn't depend on  $X$  anymore

$$p(x, \theta|X) = p(x|\theta) p(\theta|X)$$

$$p(x|X) = \int p(x|\theta) p(\theta|X) d\theta$$

This is entirely determined by the parameter  $\theta$  (i.e. by the parametric form of the pdf).

37

RWTH AACHEN UNIVERSITY

## Bayesian Learning Approach

$$p(x|X) = \int p(x|\theta)p(\theta|X)d\theta$$

$$p(\theta|X) = \frac{p(X|\theta)p(\theta)}{p(X)} = \frac{p(\theta)}{p(X)}L(\theta)$$

$$p(X) = \int p(X|\theta)p(\theta)d\theta = \int L(\theta)p(\theta)d\theta$$

- Inserting this above, we obtain

$$p(x|X) = \int \frac{p(x|\theta)L(\theta)p(\theta)}{p(X)}d\theta = \int \frac{p(x|\theta)L(\theta)p(\theta)}{\int L(\theta)p(\theta)d\theta}d\theta$$

Machine Learning Summer'16 38

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Bayesian Learning Approach

- Discussion
  - Likelihood of the parametric form  $\theta$  given the data set  $X$ .
  - Prior for the parameters  $\theta$
  - Estimate for  $x$  based on parametric form  $\theta$

$$p(x|X) = \int \frac{p(x|\theta)L(\theta)p(\theta)}{\int L(\theta)p(\theta)d\theta}d\theta$$

Normalization: integrate over all possible values of  $\theta$

- If we now plug in a (suitable) prior  $p(\theta)$ , we can estimate  $p(x|X)$  from the data set  $X$ .

Machine Learning Summer'16 39

B. Leibe

RWTH AACHEN UNIVERSITY

## Bayesian Density Estimation

- Discussion

$$p(x|X) = \int p(x|\theta)p(\theta|X)d\theta = \int \frac{p(x|\theta)L(\theta)p(\theta)}{\int L(\theta)p(\theta)d\theta}d\theta$$

- The probability  $p(\theta|X)$  makes the dependency of the estimate on the data explicit.
- If  $p(\theta|X)$  is very small everywhere, but is large for one  $\hat{\theta}$ , then
 
$$p(x|X) \approx p(x|\hat{\theta})$$

⇒ The more uncertain we are about  $\theta$ , the more we average over all parameter values.

Machine Learning Summer'16 40

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Bayesian Density Estimation

- Problem
  - In the general case, the integration over  $\theta$  is not possible (or only possible stochastically).
- Example where an analytical solution is possible
  - Normal distribution for the data,  $\sigma^2$  assumed known and fixed.
  - Estimate the distribution of the mean:
 
$$p(\mu|X) = \frac{p(X|\mu)p(\mu)}{p(X)}$$
  - Prior: We assume a Gaussian prior over  $\mu$ ,
 
$$p(\mu) = \mathcal{N}(\mu|\mu_0, \sigma_0^2)$$

Machine Learning Summer'16 41

Slide credit: Bernt Schiele B. Leibe

RWTH AACHEN UNIVERSITY

## Bayesian Learning Approach

- Sample mean:  $\bar{x} = \frac{1}{N} \sum_{n=1}^N x_n$
- Bayes estimate:
 
$$\mu_N = \frac{\sigma^2 \mu_0 + N \sigma_0^2 \bar{x}}{\sigma^2 + N \sigma_0^2}$$

$$\frac{1}{\sigma_N^2} = \frac{1}{\sigma_0^2} + \frac{N}{\sigma^2}$$
- Note:
 

	$N = 0$	$N \rightarrow \infty$
$\mu_N$	$\mu_0$	$\mu_{ML}$
$\sigma_N^2$	$\sigma_0^2$	0

Machine Learning Summer'16 42

Slide adapted from Bernt Schiele B. Leibe Image source: C. M. Bishop, 2006

RWTH AACHEN UNIVERSITY

## Summary: ML vs. Bayesian Learning

- Maximum Likelihood
  - Simple approach, often analytically possible.
  - Problem: estimation is biased, tends to overfit to the data.
    - ⇒ Often needs some correction or regularization.
  - But:
    - Approximation gets accurate for  $N \rightarrow \infty$ .
- Bayesian Learning
  - General approach, avoids the estimation bias through a prior.
  - Problems:
    - Need to choose a suitable prior (not always obvious).
    - Integral over  $\theta$  often not analytically feasible anymore.
  - But:
    - Efficient stochastic sampling techniques available (see Adv. ML).

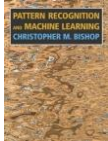
(In this lecture, we'll use both concepts wherever appropriate)

Machine Learning Summer'16 43

B. Leibe

## References and Further Reading

- More information in Bishop's book
  - Gaussian distribution and ML: Ch. 1.2.4 and 2.3.1-2.3.4.
  - Bayesian Learning: Ch. 1.2.3 and 2.3.6.
  - Nonparametric methods: Ch. 2.5.
- Additional information can be found in Duda & Hart
  - ML estimation: Ch. 3.2
  - Bayesian Learning: Ch. 3.3-3.5
  - Nonparametric methods: Ch. 4.1-4.5



Christopher M. Bishop  
Pattern Recognition and Machine Learning  
Springer, 2006



R.O. Duda, P.E. Hart, D.G. Stork  
Pattern Classification  
2<sup>nd</sup> Ed., Wiley-Interscience, 2000  
B. Leibe