

Computer Vision II - Lecture 10

Particle Filters (The Gritty Details)

27.05.2014

Bastian Leibe RWTH Aachen http://www.vision.rwth-aachen.de

leibe@vision.rwth-aachen.de



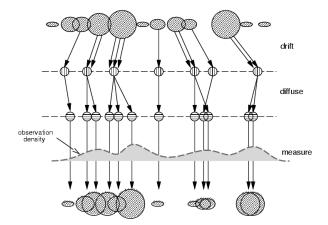
Announcement

- Problems with exam registration fixed...
 - ...for Master CS and Master SSE
 - You should now be able to register
 - I extended the registration deadline until this Friday (30.05.)
- Exchange students can register directly with us
 - If registration is not possible via ZPA
- Please let us know if problems persist.

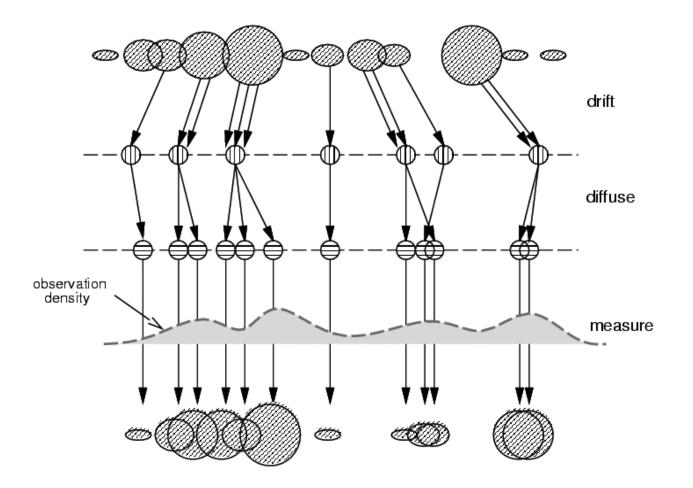


Course Outline

- Single-Object Tracking
 - Background modeling
 - > Template based tracking
 - Color based tracking
 - Contour based tracking
 - Tracking by online classification
 - Tracking-by-detection
- Bayesian Filtering
 - Kalman filters
 - Particle filters
 - Case studies
- Multi-Object Tracking
- Articulated Tracking



Today: Beyond Gaussian Error Models





Topics of This Lecture

- Recap: Extended Kalman Filter
 - Detailed algorithm
- Particle Filters: Detailed Derivation
 - Recap: Basic idea
 - > Importance Sampling
 - Sequential Importance Sampling (SIS)
 - Transitional prior
 - Resampling
 - Generic Particle Filter
 - Sampling Importance Resampling (SIR)



Recap: Kalman Filter

- Algorithm summary
 - > Assumption: linear model

$$\mathbf{x}_t = \mathbf{D}_t \mathbf{x}_{t-1} + \varepsilon_t$$
$$\mathbf{y}_t = \mathbf{M}_t \mathbf{x}_t + \delta_t$$

> Prediction step

$$egin{array}{rcl} \mathbf{x}_t^- &=& \mathbf{D}_t \mathbf{x}_{t-1}^+ \ \mathbf{\Sigma}_t^- &=& \mathbf{D}_t \mathbf{\Sigma}_{t-1}^+ \mathbf{D}_t^T + \mathbf{\Sigma}_{d_t} \end{array}$$

Correction step

$$\begin{split} \mathbf{K}_t &= \mathbf{\Sigma}_t^{-} \mathbf{M}_t^T \left(\mathbf{M}_t \mathbf{\Sigma}_t^{-} \mathbf{M}_t^T + \mathbf{\Sigma}_{m_t} \right)^{-1} \\ \mathbf{x}_t^{+} &= \mathbf{x}_t^{-} + \mathbf{K}_t \left(\mathbf{y}_t - \mathbf{M}_t \mathbf{x}_t^{-} \right) \\ \mathbf{\Sigma}_t^{+} &= \left(\mathbf{I} - \mathbf{K}_t \mathbf{M}_t \right) \mathbf{\Sigma}_t^{-} \end{split}$$

Recap: Extended Kalman Filter (EKF)

- Algorithm summary
 - Nonlinear model

$$\mathbf{x}_{t} = \mathbf{g}(\mathbf{x}_{t-1}) + \varepsilon_{t}$$
$$\mathbf{y}_{t} = \mathbf{h}(\mathbf{x}_{t}) + \delta_{t}$$

Prediction step

with the Jacobians

- $egin{array}{lll} \mathbf{x}_t^- &=& \mathbf{g}\left(\mathbf{x}_{t-1}^+
 ight) \ \mathbf{\Sigma}_t^- &=& \mathbf{G}_t\mathbf{\Sigma}_{t-1}^+\mathbf{G}_t^T + \mathbf{\Sigma}_{d_t} \end{array} \qquad \qquad \mathbf{G}_t &=& rac{\partial \mathbf{g}(\mathbf{x})}{\partial \mathbf{x}} igg|_{\mathbf{x}_t = \mathbf{x}_t^+} \end{array}$



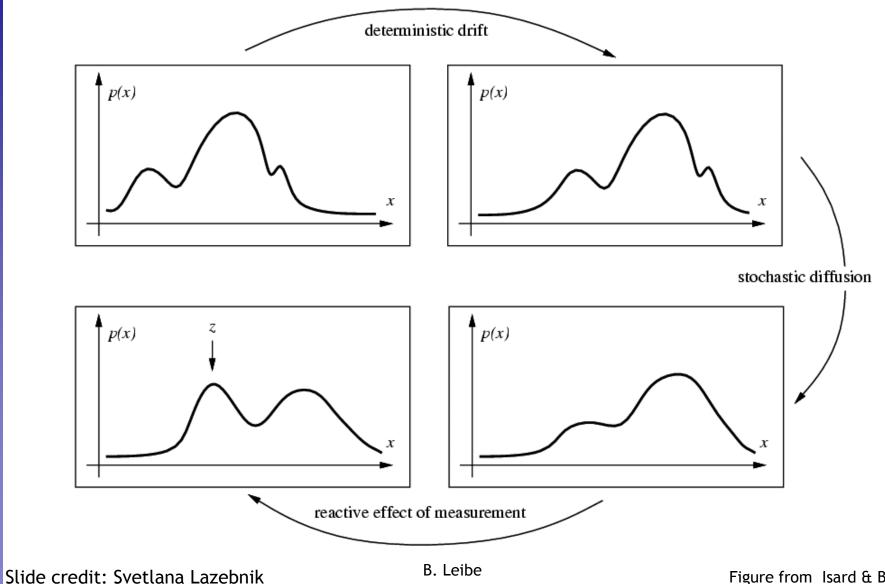
Topics of This Lecture

Recap: Extended Kalman Filter
 Detailed algorithm

• Particle Filters: Detailed Derivation

- Recap: Basic idea
- > Importance Sampling
- Sequential Importance Sampling (SIS)
- Transitional prior
- Resampling
- Generic Particle Filter
- Sampling Importance Resampling (SIR)

Recap: Propagation of General Densities

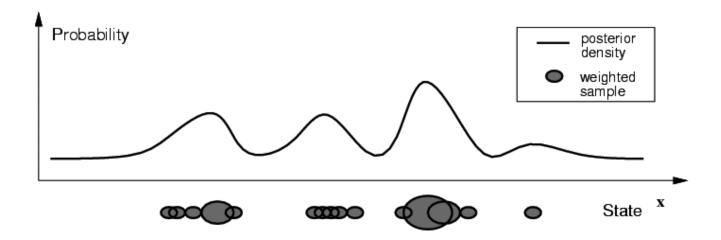


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9 Figure from Isard & Blake



Recap: Factored Sampling



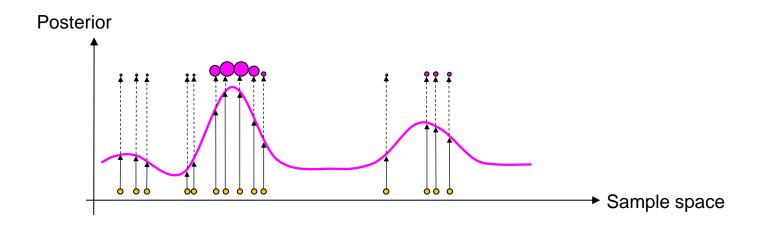
- Idea: Represent state distribution non-parametrically
 - > Prediction: Sample points from prior density for the state, P(X)
 - > Correction: Weight the samples according to P(Y|X)

$$P(X_{t} | y_{0},..., y_{t}) = \frac{P(y_{t} | X_{t})P(X_{t} | y_{0},..., y_{t-1})}{\int P(y_{t} | X_{t})P(X_{t} | y_{0},..., y_{t-1})dX_{t}}$$



Particle Filtering

- Many variations, one general concept:
 - Represent the posterior pdf by a set of randomly chosen weighted samples (particles)



- Randomly Chosen = Monte Carlo (MC)
- As the number of samples become very large the characterization becomes an equivalent representation of the true pdf.



Particle filtering

- Compared to Kalman Filters and their extensions
 - > Can represent any arbitrary distribution
 - Multimodal support
 - > Keep track of as many hypotheses as there are particles
 - Approximate representation of complex model rather than exact representation of simplified model
- The basic building-block: Importance Sampling



f(z)

Recap: Monte-Carlo Sampling

- Objective:
 - > Evaluate expectation of a function f(z)w.r.t. a probability distribution p(z).

 $\mathbb{E}[f] = \int f(\mathbf{z}) p(\mathbf{z}) d\mathbf{z}$

- Monte Carlo Sampling idea
 - > Draw L independent samples $z^{(l)}$ with l = 1, ..., L from p(z).
 - This allows the expectation to be approximated by a finite sum

p(z)

$$\hat{f} = \frac{1}{L} \sum_{l=1}^{L} f(\mathbf{z}^l)$$

As long as the samples $\mathbf{z}^{(l)}$ are drawn independently from $p(\mathbf{z})$, then $\mathbb{E}[\hat{f}] = \mathbb{E}[f]$

\Rightarrow Unbiased estimate, independent of the dimension of z!

Slide adapted from Bernt Schiele



Monte Carlo Integration

- We can use the same idea for computing integrals
 - > Assume we are trying to estimate a complicated integral of a function f over some domain D:

$$F = \int_D f(\vec{x}) d\vec{x}$$

 \succ Also assume there exists some PDF p defined over D. Then

$$F = \int_D f(\vec{x}) d\vec{x} = \int_D \frac{f(\vec{x})}{p(\vec{x})} p(\vec{x}) d\vec{x}$$

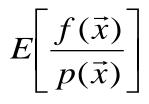
 \succ For any pdf p over D, the following holds

$$\int_{D} \frac{f(\vec{x})}{p(\vec{x})} p(\vec{x}) d\vec{x} = E\left[\frac{f(\vec{x})}{p(\vec{x})}\right], x \sim p$$



Monte Carlo Integration

- Idea (cont'd)
 - > Now, if we have i.i.d random samples x_1, \ldots, x_N sampled from p, then we can approximate the expectation



by

$$F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(\vec{x}_i)}{p(\vec{x}_i)}$$

Guaranteed by law of large numbers:

$$N \to \infty, F_N \xrightarrow{a.s} E\left[\frac{f(\vec{x})}{p(\vec{x})}\right] = F$$

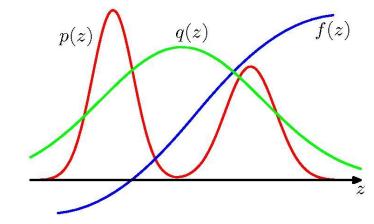
> Since it guides sampling, p is often called a proposal distribution.



Importance Sampling

• Let's consider an example

$$F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(\vec{x}_i)}{p(\vec{x}_i)}$$



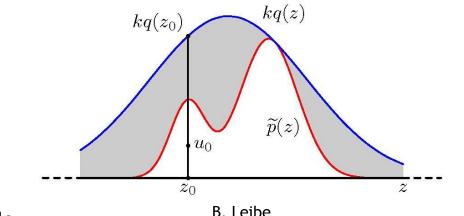
- > f/p is the importance weight of a sample.
- What can go wrong here?
- What if p(x)=0 ?
 - > If p is very small, then f/p can get arbitrarily large!
 - \Rightarrow Design p such that f/p is bounded.
 - > Effect: get more samples in "important" areas of f, i.e., where f is large.

Proposal Distributions: Other Uses

- **Similar Problem**
 - > For many distributions, sampling directly from p(z) is difficult.
 - But we can often easily *evaluate* p(z) (up to some normalization \geq factor Z_p): $p(\mathbf{z}) = \frac{1}{Z_n} \tilde{p}(\mathbf{z})$

Idea

Take some simpler distribution q(z) as proposal distribution \geq from which we can draw samples and which is non-zero.



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Recap: Importance Sampling

• Idea

- > Use a proposal distribution q(z) from which it is easy to draw samples and which is close in shape to f.
- > Express expectations in the form of a finite sum over samples $\{z^{(l)}\}\$ drawn from q(z).

$$\mathbb{E}[f] = \int f(\mathbf{z})p(\mathbf{z})d\mathbf{z} = \int f(\mathbf{z})\frac{p(\mathbf{z})}{q(\mathbf{z})}q(\mathbf{z})d\mathbf{z}$$

$$\simeq \frac{1}{L}\sum_{l=1}^{L}\frac{p(\mathbf{z}^{(l)})}{q(\mathbf{z}^{(l)})}f(\mathbf{z}^{(l)})$$
with importance weights
$$r_{l} = \frac{p(\mathbf{z}^{(l)})}{q(\mathbf{z}^{(l)})}$$

Slide credit: Bernt Schiele

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18 Image source: C.M. Bishop, 2006

f(z)

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Illustration of Importance Factors

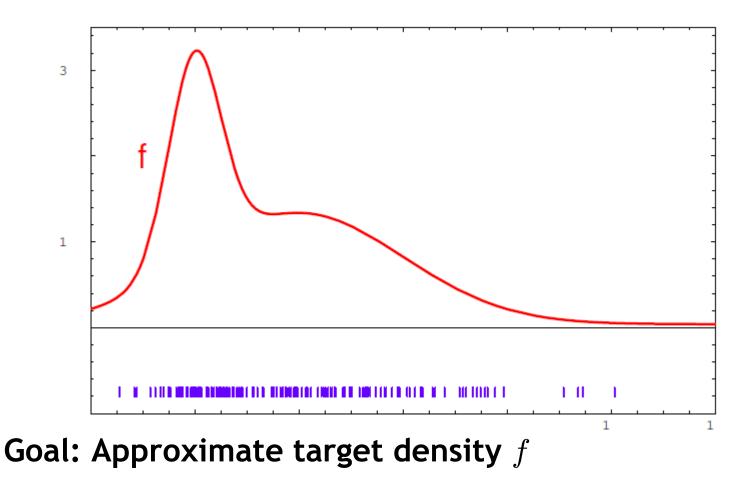
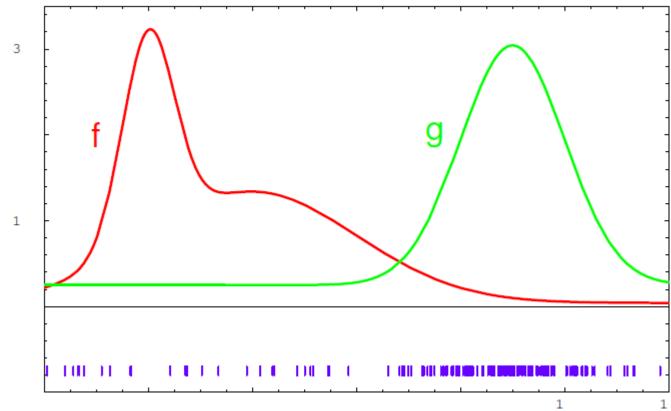




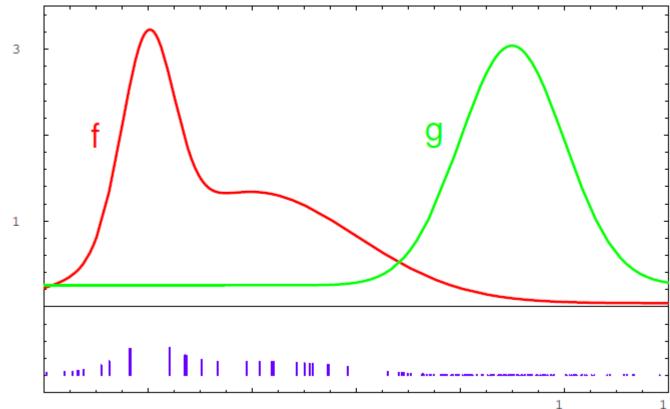
Illustration of Importance Factors



- Goal: Approximate target density f
 - > Instead of sampling from f directly, we can only sample from g.



Illustration of Importance Factors

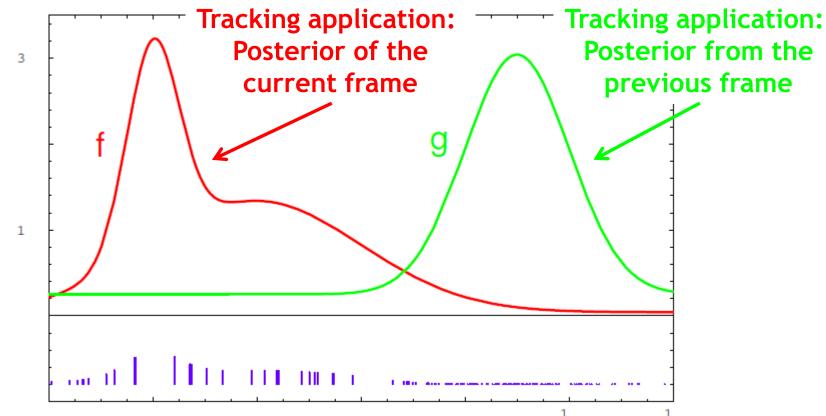


- Goal: Approximate target density f
 - > Instead of sampling from f directly, we can only sample from g.
 - > A sample of f is obtained by attaching the weight f/g to each sample \mathbf{x} .

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Interpretation for Tracking



- Goal: Approximate target density f
 - > Instead of sampling from f directly, we can only sample from g.
 - > A sample of f is obtained by attaching the weight f/g to each sample \mathbf{x} .

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Importance Sampling for Bayesian Estimation

$$\mathbb{E}[f(X)] = \int_X f(\mathbf{x}_{0:t}) p(\mathbf{x}_{0:t} | \mathbf{y}_{1:t}) d\mathbf{x}_{0:t}$$

=
$$\int_X f(\mathbf{x}_{0:t}) \frac{p(\mathbf{x}_{0:t} | \mathbf{y}_{1:t})}{q(\mathbf{x}_{0:t} | \mathbf{y}_{1:t})} q(\mathbf{x}_{0:t} | \mathbf{y}_{1:t}) d\mathbf{x}_{0:t}$$

- Applying Importance Sampling
 - Characterize the posterior pdf using a set of samples (particles) and their weights

$$\left\{\mathbf{x}_{0:t}^{i}, w_{t}^{i}\right\}_{i=1}^{N}$$

Then the joint posterior is approximated by

$$p(\mathbf{x}_{0:t}|\mathbf{y}_{1:t}) \approx \sum_{i=1}^{N} w_t^i \delta(\mathbf{x}_{0:t} - \mathbf{x}_{0:t}^i)$$

Importance Sampling for Bayesian Estimation

$$\mathbb{E}[f(X)] = \int_X f(\mathbf{x}_{0:t}) p(\mathbf{x}_{0:t} | \mathbf{y}_{1:t}) d\mathbf{x}_{0:t}$$

=
$$\int_X f(\mathbf{x}_{0:t}) \frac{p(\mathbf{x}_{0:t} | \mathbf{y}_{1:t})}{q(\mathbf{x}_{0:t} | \mathbf{y}_{1:t})} q(\mathbf{x}_{0:t} | \mathbf{y}_{1:t}) d\mathbf{x}_{0:t}$$

- Applying Importance Sampling
 - > Draw the samples from the importance density $q(\mathbf{x}_{0:t} | \mathbf{y}_{1:t})$ with importance weights $p(\mathbf{x}_{0:t} | \mathbf{y}_{1:t})$

$$w_t^i \propto rac{p(\mathbf{x}_{0:t}|\mathbf{y}_{1:t})}{q(\mathbf{x}_{0:t}|\mathbf{y}_{1:t})}$$

- Sequential update (after some calculation)
 - Particle update $\mathbf{x}_t \sim q(\mathbf{x}_t | \mathbf{x}_{t-1}^i, \mathbf{y}_t)$

- Weight update
$$w_t^i = w_{t-1}^i \frac{p(\mathbf{y}_t | \mathbf{x}_t^i) p(\mathbf{x}_t^i | \mathbf{x}_{t-1}^i)}{q(\mathbf{x}_t | \mathbf{x}_{t-1}^i, \mathbf{y}_t)}$$

Slide adapted from Michael Rubinstein

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Sequential Importance Sampling Algorithm

$$\begin{aligned} & \textbf{function} \left[\left\{ \mathbf{x}_{t}^{i}, w_{t}^{i} \right\}_{i=1}^{N} \right] = SIS \left[\left\{ \mathbf{x}_{t-1}^{i}, w_{t-1}^{i} \right\}_{i=1}^{N}, \mathbf{y}_{t} \right] \\ & \eta = 0 & \text{Initialize} \\ & \textbf{for} \quad i = 1:N \\ & \mathbf{x}_{t}^{i} \sim q(\mathbf{x}_{t} | \mathbf{x}_{t-1}^{i}, \mathbf{y}_{t}) & \text{Sample from proposal} \\ & w_{t}^{i} = w_{t-1}^{i} \frac{p(\mathbf{y}_{t} | \mathbf{x}_{t}^{i}) p(\mathbf{x}_{t}^{i} | \mathbf{x}_{t-1}^{i})}{q(\mathbf{x}_{t} | \mathbf{x}_{t-1}^{i}, \mathbf{y}_{t})} & \text{Update weights} \\ & \eta = \eta + w_{t}^{i} & \text{Update norm. factor} \\ & \textbf{end} \end{aligned}$$

for i = 1:N

$$w_t^i = w_t^i / \eta$$

end

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Slide adapted from Michael Rubinstein

pdf

Normalize weights

Sequential Importance Sampling Algorithm

$$\begin{aligned} & \text{function } \left[\left\{ \mathbf{x}_{t}^{i}, w_{t}^{i} \right\}_{i=1}^{N} \right] = SIS \left[\left\{ \mathbf{x}_{t-1}^{i}, w_{t-1}^{i} \right\}_{i=1}^{N}, \mathbf{y}_{t} \right] \\ & \eta = 0 & \text{Initialize} \\ & \text{for } i = 1:N & \text{Sample from proposal pdf} \\ & \mathbf{x}_{t}^{i} \sim q(\mathbf{x}_{t} | \mathbf{x}_{t-1}^{i}, \mathbf{y}_{t}) & \text{Sample from proposal pdf} \\ & w_{t}^{i} = w_{t-1}^{i} \frac{p(\mathbf{y}_{t} | \mathbf{x}_{t}^{i}) p(\mathbf{x}_{t}^{i} | \mathbf{x}_{t-1}^{i})}{q(\mathbf{x}_{t} | \mathbf{x}_{t-1}^{i}, \mathbf{y}_{t})} & \text{Update weights} \\ & \eta = \eta + w_{t}^{i} & \text{Update norm. factor} \\ & \text{end} & \text{for a concrete algorithm,} \\ & w_{t}^{i} = w_{t}^{i} / \eta & \text{Normalize weights} \end{aligned}$$

end



Choice of Importance Density

- Most common choice
 - > Transitional prior

$$q(\mathbf{x}_t | \mathbf{x}_{t-1}^i, \mathbf{y}_t) = p(\mathbf{x}_t | \mathbf{x}_{t-1}^i)$$

> With this choice, the weight update equation simplifies to

$$w_t^i = w_{t-1}^i \frac{p(\mathbf{y}_t | \mathbf{x}_t^i) p(\mathbf{x}_t | \mathbf{x}_{t-1}^i)}{q(\mathbf{x}_t | \mathbf{x}_{t-1}^i, \mathbf{y}_t)}$$
$$= w_{t-1}^i \frac{p(\mathbf{y}_t | \mathbf{x}_t^i) p(\mathbf{x}_t | \mathbf{x}_{t-1}^i)}{p(\mathbf{x}_t | \mathbf{x}_{t-1}^i)}$$
$$= w_{t-1}^i p(\mathbf{y}_t | \mathbf{x}_t^i)$$

SIS Algorithm with Transitional Prior

$$\begin{aligned} & \mathbf{function} \left[\left\{ \mathbf{x}_{t}^{i}, w_{t}^{i} \right\}_{i=1}^{N} \right] = SIS \left[\left\{ \mathbf{x}_{t-1}^{i}, w_{t-1}^{i} \right\}_{i=1}^{N}, \mathbf{y}_{t} \right] \\ & \eta = 0 & \text{Initialize} \\ & \text{for } i = 1:N \\ & \mathbf{x}_{t}^{i} \sim p(\mathbf{x}_{t} | \mathbf{x}_{t-1}^{i}) & \text{Sample from proposal pdf} \\ & w_{t}^{i} = w_{t-1}^{i} p(\mathbf{y}_{t} | \mathbf{x}_{t}^{i}) & \text{Update weights} \\ & \eta = \eta + w_{t}^{i} & \text{Update norm. factor} \\ & \text{end} \\ & \text{for } i = 1:N \\ & w_{t}^{i} = w_{t}^{i} / \eta & \text{Normalize weights} \\ & \text{end} \end{aligned}$$

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SIS Algorithm with Transitional Prior

$$\begin{aligned} & \mathbf{function} \left[\left\{ \mathbf{x}_{t}^{i}, w_{t}^{i} \right\}_{i=1}^{N} \right] = SIS \left[\left\{ \mathbf{x}_{t-1}^{i}, w_{t-1}^{i} \right\}_{i=1}^{N}, \mathbf{y}_{t} \right] \\ & \eta = 0 & \text{Initialize} \\ & \text{for } i = 1:N \\ & Draw \ \varepsilon_{t}^{i} \ from \ noise \ distribution \\ & \mathbf{x}_{t}^{i} = \mathbf{g} \left(\mathbf{x}_{t-1}^{i} \right) + \varepsilon_{t}^{i} & \text{Sample from proposal pdf} \\ & w_{t}^{i} = w_{t-1}^{i} p(\mathbf{y}_{t} | \mathbf{x}_{t}^{i}) & \text{Update weights} \\ & \eta = \eta + w_{t}^{i} & \text{Update norm. factor} \\ & \text{end} \\ & \text{for } i = 1:N \\ & w_{t}^{i} = w_{t}^{i} / \eta & \text{Normalize weights} \end{aligned}$$

end

Slide adapted from Michael Rubinstein

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The Degeneracy Phenomenon

- Unavoidable problem with SIS
 - After a few iterations, most particles have negligible weights.
 - Large computational effort for updating particles with very small contribution to $p(\mathbf{x}_t \mid \mathbf{y}_{1:t})$.
- Measure of degeneracy
 - Effective sample size

$$N_{eff} = \frac{1}{\sum_{i=1}^{N} (w_t^i)^2}$$

- **Uniform:**
- $N_{eff} = N$ $N_{eff} = 1$ Severe degeneracy:



Resampling

• Idea

Eliminate particles with low importance weights and increase the number of particles with high importance weight.

$$\left\{\mathbf{x}_{t}^{i}, w_{t}^{i}\right\}_{i=1}^{N} \rightarrow \left\{\mathbf{x}_{t}^{i*}, \frac{1}{N}\right\}_{i=1}^{N}$$

> The new set is generated by sampling with replacement from the discrete representation of $p(\mathbf{x}_t \mid \mathbf{y}_{1:t})$ such that

$$Pr\left\{\mathbf{x}_t^{i*} = \mathbf{x}_t^j\right\} = w_t^j$$

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Resampling

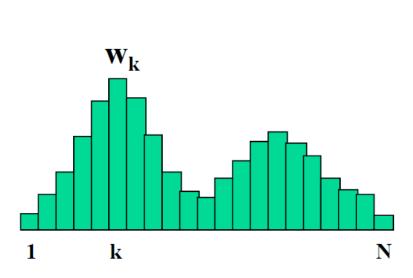
- How to do that in practice?
 - > We want to resample $\{\mathbf{x}_t^i\}_{i=1}^N$ from the discrete pdf given by the weighted samples $\{\mathbf{x}_t^i, w_t^i\}_{i=1}^N$.
 - > I.e., we want to draw N new samples $\{\mathbf{x}_t^i\}_{i=1}^N$ with replacement where the probability of drawing \mathbf{x}_t^j is given by w_t^j .
- There are many algorithms for this
 - > We will look at two simple algorithms here...

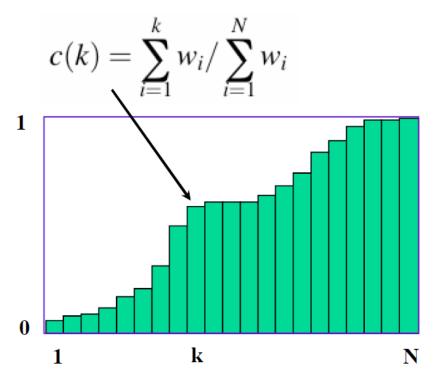


Inverse Transform Sampling

• Idea

> It is easy to sample from a discrete distribution using the cumulative distribution function $F(x) = p(X \le x)$.





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Inverse Transform Sampling

• Idea

- > It is easy to sample from a discrete distribution using the cumulative distribution function $F(x) = p(X \le x)$.
- $c(k) = \sum_{i=1}^{n} w_i / \sum_{i=1}^{n} w_i$ Procedure **1.** Generate uniform *u* in the range [0,1]. 2. Visualize a horizontal line intersecting the bars. u3. If index of intersected bar is j, output new sample \mathbf{x}_i . 1 k

Slide adapted from Robert Collins

N



More Efficient Approach

• From Arulampalam paper:

Algorithm 2: Resampling Algorithm $[\{\mathbf{x}_{k}^{j*}, w_{k}^{j}, i^{j}\}_{j=1}^{N_{s}}] = \text{RESAMPLE} [\{\mathbf{x}_{k}^{i}, w_{k}^{i}\}_{i=1}^{N_{s}}]$ • Initialize the CDF: $c_1 = 0$ • FOR i = 2: N_s - Construct CDF: $c_i = c_{i-1} + w_k^i$ END FOR • Start at the bottom of the CDF: i=1• Draw a starting point: $u_1 \sim \mathbb{V}[0, N_s^{-1}]$ • FOR j = 1: N_s - Move along the CDF: $u_j = u_1 + N_s^{-1}(j-1)$ - WHILE $u_j > c_i$ * i = i + 1- END WHILE - Assign sample: $\mathbf{x}_k^{j*} = \mathbf{x}_k^i$ - Assign weight: $w_k^j = N_s^{-1}$ - Assign parent: $i^{j} = i$ END FOR

Basic idea: choose one initial small random number; deterministically sample the rest by "crawling" up the cdf. This is $\mathcal{O}(N)$!

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Generic Particle Filter

$$\begin{aligned} \mathbf{function} & \left[\left\{ \mathbf{x}_{t}^{i}, w_{t}^{i} \right\}_{i=1}^{N} \right] = PF\left[\left\{ \mathbf{x}_{t-1}^{i}, w_{t-1}^{i} \right\}_{i=1}^{N}, \mathbf{y}_{t} \right] \\ Apply SIS filtering & \left[\left\{ \mathbf{x}_{t}^{i}, w_{t}^{i} \right\}_{i=1}^{N} \right] = SIS\left[\left\{ \mathbf{x}_{t-1}^{i}, w_{t-1}^{i} \right\}_{i=1}^{N}, \mathbf{y}_{t} \right] \\ Calculate \ N_{eff} \end{aligned}$$

if
$$N_{eff} < N_{thr}$$

$$\left[\left\{\mathbf{x}_{t}^{i}, w_{t}^{i}\right\}_{i=1}^{N}\right] = RESAMPLE\left[\left\{\mathbf{x}_{t}^{i}, w_{t}^{i}\right\}_{i=1}^{N}\right]$$
end

- We can also apply resampling selectively
 - > Only resample when it is needed, i.e., N_{eff} is too low.
 - \Rightarrow Avoids drift when there the tracked state is stationary.



Other Variant of the Algorithm

function $[\mathcal{X}_t] = SIR [\mathcal{X}_{t-1}, \mathbf{y}_t]$ $\bar{\mathcal{X}}_t = \mathcal{X}_t = \emptyset$ for i = 1:NSample $\mathbf{x}_t^i \sim p(\mathbf{x}_t | \mathbf{x}_{t-1}^i)$ $w_t^i = p(\mathbf{y}_t | \mathbf{x}_t^i)$ end for i = 1:NDraw i with probability $\propto w_t^i$ Add \mathbf{x}_t^i to \mathcal{X}_t

Initialize

Generate new samples

Update weights

Resample

end



Other Variant of the Algorithm

function $[\mathcal{X}_t] = SIR [\mathcal{X}_{t-1}, \mathbf{y}_t]$ $\mathcal{X}_t = \mathcal{X}_t = \emptyset$ for i = 1:NSample $\mathbf{x}_t^i \sim p(\mathbf{x}_t | \mathbf{x}_{t-1}^i)$ $w_t^i = p(\mathbf{y}_t | \mathbf{x}_t^i)$ end for i = 1:NDraw i with probability $\propto w_t^i$ Add \mathbf{x}_t^i to \mathcal{X}_t

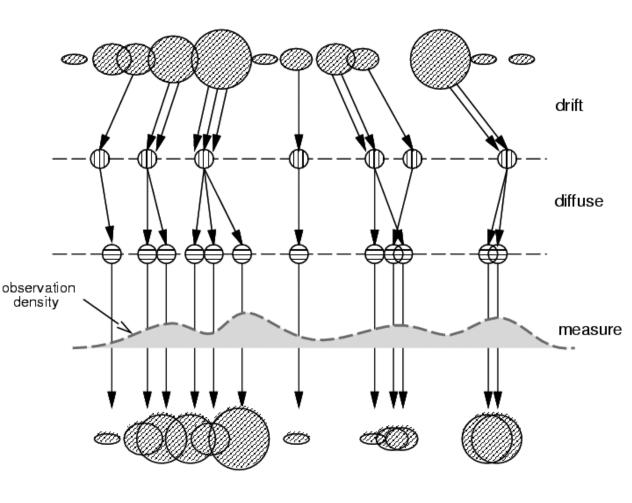
Important property:

Particles are distributed according to pdf from previous time step.

Particles are distributed according to posterior from this time step.

end

RWITHAACHEN UNIVERSITY Particle Filtering: Condensation Algorithm



Start with weighted samples from previous time step

Sample and shift according to dynamics model

Spread due to randomness; this is predicted density $p(x_t|y_{t-1})$

Weight the samples according to observation density

Arrive at corrected density estimate $p(x_t | y_t)$

M. Isard and A. Blake, <u>CONDENSATION -- conditional density propagation for</u> <u>visual tracking</u>, IJCV 29(1):5-28, 1998

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43 Figure source: M. Isard & A. Blake



Summary: Particle Filtering

• <u>Pros:</u>

- > Able to represent arbitrary densities
- Converging to true posterior even for non-Gaussian and nonlinear system
- > Efficient: particles tend to focus on regions with high probability
- Works with many different state spaces
 - E.g. articulated tracking in complicated joint angle spaces
- Many extensions available



Summary: Particle Filtering

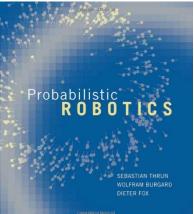
Cons / Caveats:

- #Particles is important performance factor
 - Want as few particles as possible for efficiency.
 - But need to cover state space sufficiently well.
- Worst-case complexity grows exponentially in the dimensions
- > Multimodal densities possible, but still single object
 - Interactions between multiple objects require special treatment.
 - Not handled well in the particle filtering framework (state space explosion).



References and Further Reading

- A good description of Particle Filters can be found in Ch.4.3 of the following book
 - S. Thrun, W. Burgard, D. Fox. <u>Probabilistic</u> <u>Robotics</u>. MIT Press, 2006.



- A good tutorial on Particle Filters
 - M.S. Arulampalam, S. Maskell, N. Gordon, T. Clapp. <u>A Tutorial</u> <u>on Particle Filters for Online Nonlinear/Non-Gaussian Bayesian</u> <u>Tracking</u>. In *IEEE Transactions on Signal Processing*, Vol. 50(2), pp. 174-188, 2002.
- The CONDENSATION paper
 - M. Isard and A. Blake, <u>CONDENSATION conditional density</u> propagation for visual tracking, IJCV 29(1):5-28, 1998